

# Multiple linear regression

Jean-Michel Marin

University of Montpellier  
Faculty of Sciences

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# Homoscedastic models

## Data

- ▶  $n$  individuals :  $i = 1, \dots, n$
- ▶  $y_i \in \mathbb{R}$  variable to explain (response)
- ▶  $\mathbf{x}_i = (x_{i1}, \dots, x_{ip}) \in \mathbb{R}^p$  values of  $p$  explanatory variables including the constant term constant  $x_{i1} = 1, \forall i = 1, \dots, n$

## Matrix notation

$$\mathbf{y} = (y_1, \dots, y_n) \quad ; \quad \mathbf{X} = \begin{bmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \\ \vdots \\ \mathbf{x}_n \end{bmatrix} = \begin{bmatrix} 1 & x_{12} & \dots & x_{1p} \\ 1 & x_{22} & \dots & x_{2p} \\ \vdots & \vdots & \vdots & \vdots \\ 1 & x_{n2} & \dots & x_{np} \end{bmatrix}$$

# Homoscedastic models

The multiple linear regression model is such that

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{u}$$

$\boldsymbol{\beta} = (\beta_1, \dots, \beta_p)$  is an unknown parameter

$\mathbf{u} = (u_1, \dots, u_n)$  vector of residuals

We can write,  $\forall i = 1, \dots, n$

$$y_i = \beta_1 + \beta_2 x_{i2} + \dots + \beta_p x_{ip} + u_i$$

$y_i$  is a random variable for which we have a realization

$\mathbf{x}_i$  is a fixed vector

# Homoscedastic models

## Hypotheses

- ▶ the constant regressor is present in all models
- ▶  $n > p$  more observations than variables
- ▶  $\mathbf{X}^T \mathbf{X}$  invertible, the rank of  $\mathbf{X}$  is equal to  $p$ , the model is identifiable
- ▶  $\mathbb{E}[\mathbf{u}] = \mathbf{0}_n$  residuals are centered
- ▶  $\mathbb{E}[\mathbf{u}\mathbf{u}^T] = \mathbb{V}(\mathbf{u}) = \sigma^2 \mathbf{I}_n$  residuals are decorrelated and have the same variance  $\sigma^2$

# Homoscedastic models

## Estimations of $\beta$ and $\sigma^2$

Ordinary least squares (OLS) estimator of  $\beta$

$$\hat{\beta} \in \arg \min_{\beta \in \mathbb{R}^p} \|\mathbf{y} - \mathbf{X}\beta\|^2 \left( = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \mathbf{x}_i^T \beta)^2 \right)$$

$$\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

Unbiased estimator of  $\sigma^2$

$$\hat{\sigma}^2 = \frac{\|\mathbf{y} - \mathbf{X}\hat{\beta}\|^2}{n - p} = \frac{\text{RSS}}{n - p}$$

Estimator of the standard error of  $\hat{\beta}_j$

$$\hat{\sigma}(\hat{\beta}_j) = \sqrt{\hat{\sigma}^2 (\mathbf{X}^T \mathbf{X})^{-1}_{(j,j)}}$$

# Homoscedastic models

## Geometric properties

$\chi$  vector subspace of  $\mathbb{R}^n$  generated by the columns of  $\mathbf{X}$

$\mathbf{H} = \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$  orthogonal projector on  $\chi$

$\mathbf{K} = \mathbf{I}_n - \mathbf{H}$  projector on  $\chi^\perp$

$$\hat{\mathbf{y}} = \mathbf{X} \hat{\boldsymbol{\beta}} = \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} = \mathbf{H} \mathbf{y}$$

$\hat{\mathbf{y}}$  orthogonal projection of  $\mathbf{y}$  on  $\chi$

$$\hat{\mathbf{u}} = \mathbf{y} - \hat{\mathbf{y}} = \mathbf{y} - \mathbf{H} \mathbf{y} = (\mathbf{I}_n - \mathbf{H}) \mathbf{y} = \mathbf{K} \mathbf{y}$$

$\hat{\mathbf{u}}$  orthogonal projection of  $\mathbf{y}$  on  $\chi^\perp$

$$\hat{\mathbf{y}} \perp \hat{\mathbf{u}}$$

# Homoscedastic models

## Geometric properties

According to the Pythagorean theorem

$$\|\mathbf{y} - \bar{y}\mathbf{1}_n\|^2 = \|\mathbf{y} - \hat{\mathbf{y}}\|^2 + \|\hat{\mathbf{y}} - \bar{y}\mathbf{1}_n\|^2$$

$$\boxed{\text{TSS} = \text{RSS} + \text{ESS}}$$

Coefficient of determination

$$\boxed{R^2 = \frac{\text{ESS}}{\text{TSS}} = 1 - \frac{\text{RSS}}{\text{TSS}}}$$



# Heteroscedastic models

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{u}$$

$n > p$ ,  $\mathbf{X}$  of full rank, identifiable model

$$\mathbb{E}(\mathbf{u}) = \mathbf{0}_n$$

$$\mathbb{V}(\mathbf{u}) = \mathbb{E}[\mathbf{u}\mathbf{u}^T] = \sigma^2\boldsymbol{\Omega} \text{ and } \boldsymbol{\Omega} \neq \mathbf{I}_n$$

$$\mathbb{C}(u_i, u_j) = \sigma^2\boldsymbol{\Omega}_{ij} \quad \text{heteroscedasticity and error correlation}$$

# Heteroscedastic models

It is assumed that  $\Omega$  is known

Let  $R$  be such that  $\Omega^{-1} = R^T R$

Set  $\mathbf{y}^* = R\mathbf{y}$ ,  $\mathbf{X}^* = R\mathbf{X}$  and  $\mathbf{u}^* = R\mathbf{u}$ .

$$\mathbf{y}^* = R\mathbf{y} = R(\mathbf{X}\boldsymbol{\beta} + \mathbf{u}) = R\mathbf{X}\boldsymbol{\beta} + R\mathbf{u} = \mathbf{X}^*\boldsymbol{\beta} + \mathbf{u}^*$$

$$\mathbb{E}(\mathbf{u}^*) = \mathbf{0}_n \quad ; \quad \mathbb{V}(\mathbf{u}^*) = R\mathbb{V}(\mathbf{u})R^T = \sigma^2 I_n$$

We have decorrelated the residuals and returned to the case of classical linear regression (homoscedastic)

# Heteroscedastic models

We can calculate the Ordinary Least Squares (OLS) estimator for the homoscedastic model

$$\mathbf{y}^* = \mathbf{X}^* \boldsymbol{\beta} + \mathbf{u}^*$$

$$\hat{\boldsymbol{\beta}}^* \in \arg \min_{\boldsymbol{\beta} \in \mathbb{R}^p} \|\mathbf{y}^* - \mathbf{X}^* \boldsymbol{\beta}\|^2$$

$$\hat{\boldsymbol{\beta}}^* = ((\mathbf{X}^*)^T \mathbf{X}^*)^{-1} (\mathbf{X}^*)^T \mathbf{y}^* = (\mathbf{X}^T \boldsymbol{\Omega}^{-1} \mathbf{X})^{-1} \mathbf{X}^T \boldsymbol{\Omega}^{-1} \mathbf{y}$$

This is the Generalized Least Squares (GLS) estimator of  $\boldsymbol{\beta}$ .

# Heteroscedastic models

It's difficult to estimate  $\Omega$  without additional constraints.  $\Omega$  of particular parametric forms

## General estimation method

If we have an estimator  $\hat{\Omega}(\beta)$  of  $\Omega$ , we can iterate until stabilization. we can iterate to stabilize

- ▶ estimate  $\beta$  with

$$\hat{\beta}^{(t)} = \left( \mathbf{X}^T \left[ \hat{\Omega} \left( \hat{\beta}^{(t-1)} \right) \right]^{-1} \mathbf{X} \right)^{-1} \mathbf{X}^T \left[ \hat{\Omega} \left( \hat{\beta}^{(t-1)} \right) \right]^{-1} \mathbf{y}$$

- ▶ estimate  $\Omega$  with

$$\hat{\Omega} \left( \hat{\beta}^{(t)} \right)$$