

Ex. 1

$$f(\kappa|\theta) = C_m^\kappa \theta^\kappa (1-\theta)^{m-\kappa} \quad \left. \begin{array}{l} \kappa \in \{0, \dots, m\} \\ \ell(\theta|\kappa) = C_m^\kappa \theta^\kappa (1-\theta)^{m-\kappa} \\ \kappa \text{ fixé (observé)} \\ \theta \text{ inconnue} \end{array} \right\}$$

$\kappa \in \{0, \dots, m\}$
 θ fixé
 κ inconnue

$$\pi(\theta) \longrightarrow \pi(\theta|\kappa) \longrightarrow \hat{\theta}$$

$$\triangleq \kappa \in \{0, \dots, m\}$$

$$\log(\ell(\theta|\kappa)) = \log(C_m^\kappa) + \kappa \log(\theta) + (m-\kappa) \log(1-\theta)$$

$$\frac{d \log \ell(\theta|\kappa)}{d\theta} = \left(\frac{\kappa}{\theta} \right) - \left(\frac{m-\kappa}{1-\theta} \right)$$

$$\frac{d^2 \log \ell(\theta|\kappa)}{(d\theta)^2} = -\frac{\kappa}{\theta^2} - \frac{(m-\kappa)}{(1-\theta)^2}$$

$$I_\kappa(\theta) = -E \left[-\frac{x}{\theta^2} - \frac{(m-x)}{(1-\theta)^2} \right]$$

$$I_\kappa(\theta) = \frac{m}{\theta} + \frac{m}{(1-\theta)} = \frac{m}{\theta(1-\theta)}$$

$$\pi_1(\theta) \propto \sqrt{|\Gamma_1(\theta)|} \pi(\theta) \quad [0,1]$$

$$\pi_2(\theta) \propto \theta^{-1/2} (1-\theta)^{1/2} \pi(\theta) \quad [0,1]$$

$\theta \sim \text{Beta}(1/2, 1/2)$

loi a priori de Jeffreys

$$\pi_1(\theta | \kappa) \propto l(\theta | \kappa) \pi_1(\theta)$$

$$\pi_2(\theta | \kappa) \propto \theta^\kappa (1-\theta)^{m-\kappa} \theta^{-1/2} (1-\theta)^{1/2} \pi(\theta) \quad [0,1]$$

$$\pi_2(\theta | \kappa) \propto \theta^{\kappa-1/2} (1-\theta)^{m-\kappa-1/2} \pi(\theta) \quad [0,1]$$

$$\theta | \kappa \sim \text{Beta} \left(\kappa + \frac{1}{2}, m - \kappa + \frac{1}{2} \right)$$

Estimateur bayésien

Choisissons l'espérance de la loi a posteriori

$$E(\theta | \kappa) = \frac{\kappa + 1/2}{\kappa + 1/2 + m - \kappa + 1/2}$$

$$\hat{\theta} = \frac{E \pi V}{E \pi V}$$

$$E(\theta | \kappa) = \frac{(\kappa + 1/2)}{(m + 1)} \neq \frac{\kappa}{m}$$

$$\leq \Pi_2(\theta|\kappa) \propto p(\theta|\kappa) \Pi_2(\theta)$$

$$\Rightarrow \Pi_2(\theta|\kappa) \propto \theta^\kappa (1-\theta)^{m-\kappa} \Pi_1(\theta)$$

[0,1]

$$\Rightarrow \theta|\kappa \sim \text{Beta}(\kappa+1, m-\kappa+1)$$

$$\Rightarrow \hat{\theta}_2 = \left[\frac{\kappa+1}{\kappa+1+m-\kappa+1} \right] = \left[\frac{\kappa+1}{m+2} \right]$$

$$\neq \left[\hat{\theta}_2 = \left[\frac{\kappa+1/2}{m+1} \right] \right] \neq \left[\hat{\theta}^{MV} = \frac{\kappa}{m} \right]$$

Ex. 6

$$\underline{\kappa} = (\kappa_1, \dots, \kappa_m)$$

$$\forall i \in \{1, \dots, m\} \quad \kappa_i \in \mathbb{R}_+^*$$

$$p(\theta|\underline{\kappa}) \propto \prod_{i=1}^m \exp\left[-\frac{1}{2} [\log|\kappa_i| - \theta]^2\right]$$

$$p(\theta|\underline{\kappa}) \propto \left\{ \exp\left[-\frac{1}{2} \sum_{i=1}^m (\log|\kappa_i| - \theta)^2\right] \right\}$$

$$\Pi(\theta) \propto \exp\left(-\frac{\theta^2}{2}\right)$$

$$\Pi(\theta|\underline{\kappa}) \propto \left[p(\theta|\underline{\kappa}) \Pi(\theta) \right]$$

$$\Pi(\theta|\underline{\kappa}) \propto \exp\left[-\frac{1}{2} \sum_{i=1}^m (\log|\kappa_i| - \theta)^2 - \frac{\theta^2}{2}\right]$$

$$\Pi(\theta|\underline{\kappa}) \propto \exp\left[-\frac{1}{2} \sum_{i=1}^m (\theta^2 - 2\theta \log|\kappa_i|) - \frac{\theta^2}{2}\right]$$

$$\pi(\theta|\underline{X}) \propto \exp\left[-\frac{1}{2} \left((n+1)\theta^2 - 2\theta \sum_{i=1}^n \log(X_i) \right)\right]$$

$$\pi(\theta|\underline{X}) \propto \exp\left[-\frac{(n+1)}{2} \left(\theta^2 - \frac{2}{n+1} \theta \sum_{i=1}^n \log(X_i) \right)\right]$$

$$\pi(\theta|\underline{X}) \propto \exp\left[-\frac{(n+1)}{2} \left(\theta - \frac{\sum_{i=1}^n \log(X_i)}{n+1} \right)^2\right]$$

$$\Rightarrow \theta|\underline{X} \sim \mathcal{N}\left(\frac{\sum_{i=1}^n \log(X_i)}{n+1}, \frac{1}{n+1}\right)$$

$$\Rightarrow \hat{\theta} = \frac{\sum_{i=1}^n \log(X_i)}{n+1}$$

$$L(\theta|\underline{X}) = (\theta - \mu)^2$$

$$E_{\theta|\underline{X}}((\theta - \mu)^2)$$

$$= E_{\theta|\underline{X}}\left[\theta^2 - 2\theta\mu + \mu^2\right]$$

$$= E_{\theta|\underline{X}}(\theta^2) - 2\mu E_{\theta|\underline{X}}(\theta) + \mu^2$$

On cherche μ^* qui minimise $h(\mu)$

$$h(r) = \mathbb{E}_{\Theta|r}(\Theta^2) - 2r \mathbb{E}_{\Theta|r}(\Theta) + r^2$$

$$h'(r) = -2 \mathbb{E}_{\Theta|r}(\Theta) + 2r$$

≥ 0 $h(\cdot)$ strictly
convex

$$r^* \text{ Tg: } -2 \mathbb{E}_{\Theta|r^*}(\Theta) + 2r^* = 0$$

$$\Leftrightarrow r^* = \mathbb{E}_{\Theta|r^*}(\Theta)$$

Esperanza de Θ en
el posterior

$$\angle(\Theta, r) = |\Theta - r|$$

r^* median de Θ en
posterior

Ex. 3

$$\theta \in \{0, 2\} \quad \mathbb{P}(\theta=0) = \mathbb{P}(\theta=2) = \frac{1}{2}$$

$$z = (z_1, z_2) \mid \begin{array}{l} z_1 \mid \theta \sim \mathcal{N}(\theta, 1) \\ z_2 \mid \theta \sim \mathcal{N}(\theta, 1) \end{array} \mid \underline{z_1 \perp z_2 \mid \theta}$$

$$\ell(\theta | z) \propto \left[\exp\left(-\frac{1}{2}(z_1 - \theta)^2\right) \exp\left(-\frac{1}{2}(z_2 - \theta)^2\right) \right]$$

$$\mathbb{P}(\theta=0 | z) \propto \ell(0 | z) \mathbb{P}(\theta=0)$$
$$\mathbb{P}(\theta=2 | z) \propto \ell(2 | z) \mathbb{P}(\theta=2)$$

$$\mathbb{P}(\theta=0 | z) \propto e^{-\frac{z_1^2}{2} - \frac{z_2^2}{2}} \left(\frac{1}{2}\right)$$

$$\mathbb{P}(\theta=2 | z) \propto e^{-\frac{(z_1-2)^2}{2} - \frac{(z_2-2)^2}{2}} \left(\frac{1}{2}\right)$$

$$\Rightarrow \mathbb{P}(\theta=0 | z) = \frac{e^{-\frac{z_1^2}{2} - \frac{z_2^2}{2}} \cancel{\left(\frac{1}{2}\right)}}{\cancel{\left(\frac{1}{2}\right)} e^{-\frac{z_1^2}{2} - \frac{z_2^2}{2}} + e^{-\frac{(z_1-2)^2}{2} - \frac{(z_2-2)^2}{2}} \cancel{\left(\frac{1}{2}\right)}}$$

$$\Rightarrow \mathbb{P}(\theta=2 | z) = \frac{\cancel{\left(\frac{1}{2}\right)} e^{-\frac{(z_1-2)^2}{2} - \frac{(z_2-2)^2}{2}}}{\cancel{\left(\frac{1}{2}\right)} e^{-\frac{z_1^2}{2} - \frac{z_2^2}{2}} + e^{-\frac{(z_1-2)^2}{2} - \frac{(z_2-2)^2}{2}} \cancel{\left(\frac{1}{2}\right)}}$$

$$\mathbb{P}(\theta=0 | z) + \mathbb{P}(\theta=2 | z) = 1$$